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Keywords

International Monetary shocks, Japanese economy, Oil/commodity prices, SVEC models

JEL Classification

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Not all international monetary shocks are alike for the Japanese economy

1.Introduction

In the intertemporal models of Svensson and Van Wijnbergen (1989) and Obstfeld and Rogoff (1995) monetary expansion within a large economy decreases real interest rates around the world and promotes aggregate demand worldwide. Work on exploring the impact of monetary expansion in a large economy on other economies has naturally focused on the role of the U.S. in the international transmission of monetary shocks. Evidence for this mechanism and that monetary changes in the U.S. have implications for other economies has been demonstrated in the literature on the diffusion of monetary shocks.¹

In this paper we examine the influence of monetary aggregates shocks in the U.S., China and the Euro area on Japan. The subject is important because China is now a large economy and the transmission channels for the influence of monetary shocks in China on other economies is likely to be different from the mechanism outlined for the influence of the U.S. on other economies. The topic is also valuable because the potential for big effects on income through a fall in Japan's interest rate caused by positive monetary shocks in a large open economy, such as the U.S. or the Euro area, are precluded, because Japan's interest rate is already at low levels. Examination of the effect of changes in international monetary aggregates on Japan's economy is informative since over the period of analysis these variables capture quantitative easing, something that interest rates, already at low levels, are unable to do.²

¹ A number of papers have examined the effect of U.S. monetary policy shocks on other economies. Kim (2001) and Canova (2005) find that monetary expansion in the U.S. causes economic expansion in the non-U.S. G-6 and in Latin America by lowering interest rates across these economies. In examining the role monetary policy of the non-U.S. G7 economies Grilli and Roubini (1995) demonstrate that it is important to control for U.S. monetary policy. Di Giovanni and Shambaugh (2008) find that high interest rates in a major country depress the domestic economy of a country that ties its exchange rate to that of the major country.

² Quantitative easing (QE) has been applied in Japan before and since the global financial crisis in 2007. Since the global financial crisis, quantitative easing has been used by the U.S. and the Euro area.

Over the period of this study, 1999:1-2012:12, up until 2005 China operated a dollar peg and since 2005 the renminbi has floated in a narrow margin around a fixed base rate defined with reference to a basket of major currencies with undisclosed weights.³ Throughout the period China has had extensive capital controls in place. In the Mundell–Fleming model with imperfect capital mobility, sterilization actions under a fixed exchange rate permit an independent monetary policy for as long as foreign exchange reserves permit.⁴

It is found that over 1999:1-2012:12 China's monetary expansion has significant effects on Japan's economy that are quite different from those of the U.S. and Euro area. The influence of China's monetary expansion on Japan is through the effect of China's growth on oil and world commodity prices, China's increased demand for imports, and China's exchange rate policy. The significant rise in commodity prices in response to monetary expansion in China is reflected in significantly higher inflation in Japan. China's monetary expansion is associated with significant increases in Japan's industrial production. These expansionary effects on Japan are consistent with China's monetary expansion accompanying decreases in the trade-weighted yen and increases in Japan's real exports.

In contrast, U.S. monetary expansion significantly reduces Japan's short-term interest rate (as hypothesized in the literature) and lowers Japan's exports and industrial production. Monetary expansion in the Euro area does not significantly affect Japan's exports, industrial production and interest rate. Japan's real trade balance improves with monetary expansion in China and worsens with monetary expansion in the U.S. A structural vector error correction model is used to examine international monetary transmission to Japan from U.S., China and

³ In the Mundell–Fleming model with perfect capital movement it is not possible to simultaneously have a fixed exchange rate and an autonomous monetary policy, consequently. A fixed exchange rate links domestic to foreign interest rates with trade in asset markets. A monetary expansion starts a temporary fall in the domestic interest rate relative to foreign interest rates, with the resulting loss of foreign exchange reserves leading to a reversal in the fall of the domestic interest rate.

⁴ Goodfriend and Prasad (2007) reason that capital controls provide room for monetary policy independence in China. During the fixed exchange rate period from 1998 to 2005, Sun (2009) finds that China operated an independent monetary policy. Ma and McCauley (2011) find that over 2006 to 2008 the renminbi strengthened slightly within a narrow band against a basket of major currencies.

the Euro area. Results are robust across a number of differently specified structural vector error correction models. Results are robust to changes in contemporaneous restrictions for the effect of international monetary variables, the interaction of foreign and domestic variables and to factor augmented VAR to identify monetary shocks.

In the next section recent developments in monetary aggregates and Japanese trade are briefly discussed. The methodology and the model for the study of the international monetary effects on Japan are presented in section 3. The empirical results are discussed in section 4 and checks for the robustness of results are presented in section 5. The results are discussed in Section 6. Section 7 concludes.

2. Background on the G4

The growing importance of China's money supply for global liquidity is illustrated in Figure 1. In Figure 1 the log of M2 money supplies expressed in U.S. dollars in China, U.S., Euro area, and Japan (the G4) over 1996:01-2012:12 are presented. By 2012 M2 in China exceeds that in the U.S., the Euro area, and in Japan. China's nominal M2 (in USD) increased on average by 19.5% per year from 1996 to 2011. The behaviour of China's nominal GDP is also strongly upward over the period, increasing on average (in U.S. dollars) by 15% per year (not shown in Figure 1).

Developments in trade between the U.S., China and the Euro area on the one hand with Japan on the other are presented in Figure 2, 3 and 4. The dramatic increase in trade between China and Japan, in absolute terms and relative Japan's total trade with all nations is illustrated in Figures 2 through 4. In Figure 2, the ratio of exports plus imports between China and Japan to total exports plus imports for Japan rises from 8.5% in 1999 to 20% in 2012. Over the same period, the ratio of exports plus imports between the U.S. (Euro area) and Japan to total exports plus imports for Japan falls from 27% (16%) in 1999 to 13% (9%) in 2012. The rise in China's influence in trade is particularly marked with regard to Japan's

imports. In Figure 4, Japan's imports from China, the U.S., and the Euro area relative to total Japanese imports were 21.5%, 8.5%, and 9%, respectively, in 2012, compared to 13%, 22%, and 13%, respectively, in 1999. As a market for Japanese goods China has also become much more important, with China absorbing more Japanese exports than the U.S. and far more than the Euro area over 2009-2012.

Indices of the nominal trade-weighted index for the yen and renminbi/dollar and renminbi/yen exchange rates are shown in Figure 5 with all indices being set at 100 at the start of the sample in January 1999. Prior to 2005 the renminbi was pegged to the U.S. dollar. Since July 21, 2005 the value of renminbi has been determined with regard to a basket of currencies among which the dollar has been pre-eminent. Indeed, it can be seen in Figure 5 that over an extended period from 2008 to 2010 the renminbi/dollar rate didn't vary (from 6.8 renminbi/dollar). Cai et al. (2012) and Fang et al. (2012) show that post 2005, consideration of the parity of the exchange rate of the renminbi against the U.S. dollar dominates the influence of the other currencies in determining the overall value of the renminbi (the other currencies include the Euro, Japanese yen and South Korean won).

In Figure 5 there is a tendency for the renminbi/dollar exchange rate to fall since 2005. as the U.S. dollar weakened against other currencies. The renminbi/yen exchange rate has risen and fallen over 1999-2012, with an index value in December 2012 of 97.925 just over 2% below the value in January 1999. Despite this small adjustment of the renminbi/yen exchange rate, the nominal trade-weighted index for the yen has risen over 1999-2012 with an index value of 121.285 in December 2012. Thus, by the end of the 1999-2012 period the yen has risen on a trade-weighted basis, but has fallen slightly against the renminbi. In line with the rise in the trade-weighted value of the yen, by the end of 1999-2012 period the U.S. dollar has weakened against the yen. The dollar/yen index (not shown) stands at 130.072 in December 2012 (compared to 100 in January 1999).

In models of the international transmission of monetary shocks discussion of effects on interest rates plays a large role. The Federal Funds rate for the U.S. and the central bank discount rates for China, Japan, and the Euro area are reported in Figure 6. The discount rate in Japan is at extremely low levels over 1999-2012. The Federal Funds rate for the U.S. has been close to zero since late 2008.

3. Methodology

The impact of a monetary expansion in the U.S., Euro area and China on the Japanese economy is analysed by constructing a structural vector error correction model (SVEC). Building on Kim and Roubini (2000), we use a mix of stationary and non-stationary variables to incorporate the co-integration relationship among non-stationary variables. The selection of variables and identification restrictions in the model will be based on Sims and Zha (1995) and Kim and Roubini (2000) to the extent possible given that the effects of U.S., Euro area and China monetary policy shocks on the Japanese economy are the focus of the analysis.

3.1. The variables and data

The variables in the basic model are introduced here. The international variables are the United States (U.S.) monetary aggregate M2 in U.S. dollars ($US M2_t$), Euro area monetary aggregate M2 in U.S. dollars ($EU M2_t$), the Chinese monetary aggregate M2 in U.S. dollars ($China M2_t$), and oil prices in U.S. dollars (OP_t). The domestic variables for Japan are the short term interest rate ($Japan IR_t$), the Japanese monetary aggregate M2 in U.S. dollars ($Japan M2_t$), the Japanese consumer price index ($Japan CPI_t$), the Japanese industrial production ($Japan IP_t$) and the Japanese nominal trade-weighted index (TWI_t).

M2 is chosen as measure of monetary aggregate since it is the broadest monetary aggregate available for these four largest economies for the full period 1999:1 to 2011:12 (e.g. Chinese M3 data is not available). Kozłuk and Mehrotra (2009) and Johansson (2012)

use M2 as the measure of China's monetary policy. Koz'luk and Mehrotra (2009) and Johansson (2012) find that China's monetary expansion significantly impacts real activity equity markets, respectively, in East and Southeast Asian countries. Fan et al. (2011) report that M2 is the key variable for the conduct of monetary policy in China.

The monetary aggregates over the period of analysis capture quantitative easing. Given that interest rates are at low levels, central banks have introduced quantitative easing (increasing the monetary base and money supply by buying securities) by central banks in an attempt to increase lending and liquidity.⁵ A short-term interest rate and M2 is included among the variables for Japan. Nakashima and Saito (2012) examine alternative functional forms for money demand in Japan and find that an interest rate semi elasticity function form is stable for demand for M1 and for M2 plus certificates of deposit.⁶

The data are monthly from 1999:1 to 2011:12. The sample period is determined by the availability of M2 data for the Euro area with the creation of the European Central Bank and by the observation that the People's Bank of China started concentrating on balance sheet adjustment for the conduct of monetary policy in 1998 (Johansson (2012)). G4 monetary aggregates and industrial production indexes, Japan's CPI and short-term interest rate (discount rate), and oil price (West Texas Intermediate crude oil price) data are from the Federal Reserve of St. Louis (FRED). An all commodity price index (U.S. dollar index) is from the World Bank and an index for all energy prices (U.S. dollar index) is from the International Monetary Fund. Japan's nominal and real trade weighted yen exchange rate,

⁵ Quantitative easing has been used because the official nominal short-term interests used for monetary policy are low and extraordinary actions are believed to be called for. The U.S. Federal Reserve announced QE1 in late November 2008, QE2 in November 2010, and QE3 on 13 September 2012. The European Central Bank has been buying covered bonds, a form of corporate debt. The Bank of Japan announced enormous expansions its asset purchase program in October 2011 and on April 4, 2013. The latter event is likely to lead to a doubling of the money supply.

⁶ Nagayasu (2007) uses reserve money as the measure of monetary policy since interest rates have fallen to low levels and finds that monetary expansion depreciates the yen but that this does not seem to expand real growth. Fujiwara (2007) investigates the effect of real balances (real M2 plus certificates of deposit) on real activity on Japan, given that interest rates are at extremely low levels, and finds that money has only a small direct effect in Japan. Hosono (2006) examines in detail the bank balance sheet mechanism by which monetary policy is transmitted in Japan.

nominal and real exports and imports and the real trade balance are from Central Bank of Japan.⁷ Monetary aggregates M2 are measured in nominal terms in line with Kim and Roubini (2000) and Koray and McMillin (1999). U.S. dollar amounts are used because the global oil price and commodity price index are U.S. dollar indices.

3.2. Unit root and Cointegration

The assumption of the VAR/VEC model requires that all variables in the model must be stationary, or that the linear combinations of non-stationary but co-integrated variables must be stationary. In Table 1 all variables in the model are shown to be first difference stationary.⁸ For robustness of the analysis, results in Table 1 are presented for both the Augmented Dickey-Fuller (ADF) test, where the null hypothesis is that the variable has a unit root, and the Kwiatkowski–Phillips–Schmidt–Shin (KPSS) test, where the inverse null hypothesis is tested.⁹ In line with the quantity theory of money, it is standard to test for cointegration relationships among the (log of) the variables prices, money and output. Swanson (1998), Bachmeier and Swanson (2005) and Garratt et. al. (2009) find evidence of this long run relationship.

In table 2, cointegration among the log of Japanese prices (CPI), money (M2) and income (industrial production) is tested using the Johansen cointegration test under different assumptions such as the inclusion of linear and quadratic trend and intercept. Results consistently show that there is a cointegration vector among these variables during the period 1999:1 to 2012:12. Consequently, a SVEC model using an error correction term (ECT) among the log of prices, money and output is preferred, allowing the short and long term properties of the data to be captured.

⁷ The yen's trade-weighted index is calculated using the weighted geometric average of the yen's exchange rates against 15 major currencies. The weights are based on the average of Japan's relative exports to trading partners.

⁸ Note that Japanese interest rate is not included in the table because the interest rate is already expressed in percentage. Interest rate is widely used in the macroeconomic literature without further transformations.

⁹ Results are also supported by the Dickey Fuller GLS (DF-GLS) and the Phillip-Perron (PP) which are available upon request.

3.3. Model and identification strategy

The SVEC model can be expressed as:

$$B_0 X_t = \beta + \sum_{i=1}^j B_i X_{t-i} + \lambda ECT_{t-1} + \sum_{i=1}^j \rho_i Z_{t-i} + \varepsilon_t \quad (1)$$

where j is the optimal lag length, determined by the Schwarz information criterion (SIC) in this case two lags are selected. X_t is vector of endogenous variables. Z_t is a vector of the country-specific exogenous variables. We use the industrial production and the interest rate of China, of the U.S. and of the Euro area as exogenous variables. The exogenous variables are included to isolate the effect of foreign monetary shocks from real activity and to distinguish between foreign money supply and demand. The B_i , λ , and ρ_i are matrices of coefficients to be estimated. The ε_t is the vector of structural changes, which is serially and mutually independent.

The vector X_t can be expressed as:

$$X_t = \begin{bmatrix} \Delta \log(US M2_t), \Delta \log(EU M2_t), \Delta \log(China M2_t), Japan IR_t, \Delta \log(Japan M2_t), \\ \Delta \log(Japan CPI_t), \Delta \log(Japan IP_t), \Delta \log(OP_t), \Delta \log(TWI_t) \end{bmatrix} \quad (2)$$

And the error correction term (ECT_t):

$$ECT_t = \log(Japan CPI_t) - \alpha - \psi \log(Japan IP_t) - \gamma \log(Japan M2_t) \quad (3)$$

$$Z_t = [US IR_t, EU IR_t, CH IR_t, \Delta \log(US IP_t), \Delta \log(EU IP_t), \Delta \log(CH IP_t)] \quad (4)$$

Non-recursive contemporaneous restrictions are preferred in this study in line with Bernanke (1986), Sims and Zha (1995), Kim and Roubini (2000) and Kim (2001). In particular our identification scheme is based on Kim and Roubini (2000), to the extent possible. Restrictions are only imposed in the contemporaneous matrix $B_0 X_t$ as:

$$B_0 X_t = \begin{bmatrix} 1 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 1 & 0 & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 1 & -a_{34} & 0 & 0 & -a_{37} & -a_{38} \\ 0 & 0 & 0 & -a_{43} & 1 & -a_{45} & -a_{46} & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & 1 & -a_{56} & -a_{57} & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 & 1 & -a_{67} & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & -a_{83} & -a_{84} & -a_{85} & -a_{86} & -a_{87} & 1 \end{bmatrix} \begin{pmatrix} \Delta \log(US M2_t) \\ \Delta \log(EU M2_t) \\ \Delta \log(China M2_t) \\ Japan IR_t \\ \Delta \log(Japan M2_t) \\ \Delta \log(Japan CPI_t) \\ \Delta \log(Japan IP_t) \\ \Delta \log(OP_t) \\ \Delta \log(TWI_t) \end{pmatrix} \quad (5)$$

In this system, the three international monetary aggregates, the U.S, Euro area and Chinese M2 and oil prices are treated as contemporaneously exogenous assuming some information delay from international to domestic variables. Consistent with Kim and Roubini (2000), and following Sims and Zha (1995), the central bank reaction function (the fourth equation in system 4) only responds contemporaneously to domestic monetary aggregates, nominal exchange rate and oil prices as information regarding other variables are not available within a month.

In the fifth equation (*Japan M2_t*), restrictions are in line with Kim (1999 and 2001) and Kim and Roubini (2000), implying that demand for real money depends on both real income and nominal interest rate. Restrictions in both the inflation and output equations (sixth and seventh equation respectively) are standard in the economic literature assuming that oil prices affect these variables in the same period, as oil is crucial inputs for many sectors.¹⁰ Kim and Roubini (2000), Kim (1999) and Kim (2001) treat oil prices and/or commodity prices as contemporaneously exogenous. The exchange rate equation is affected contemporaneously by all other variables as international exchange rate operators may arbitrage daily with all available information.

3.3. Test of stability, Heterosketasticity and stability condition

¹⁰ These restrictions have been also been used by Gordon and Leeper (1994), Sims and Zha (2006), Chistiano *et al.* (1999) and Kim (1999).

To test for autocorrelation the residual serial correlation LM test is carried out and p-values results estimated. The null hypothesis of no serial correlation cannot be rejected at conventional levels for the model set out in equations (1) to (4) and for the alternatives specifications introduced later. Therefore autocorrelation does not seem to be a problem in these models. Results are also confirmed by both the VAR residual cross-correlation (correlograms) and by VAR residual Portmanteau test for autocorrelation.

To test for heteroskedasticity, the VAR residual heteroskedasticity test is performed. In this test, all possible combinations of error term products are used as dependent variables. The null hypothesis of no heteroskedasticity cannot be rejected at 5% level for the joint hypothesis of all combinations for our models. The White heteroskedasticity test either using no cross terms and cross terms also confirms the previous results. Therefore, the hypothesis of heteroskedasticity in our models is discounted.

The stability condition for the VEC model, with r co-integrating equation and k endogenous variables, requires that at most $k-r$ roots should be equal to unity, while the other roots must lie inside the unit root circle.¹¹ Consequently, the inverse roots of the characteristic AR polynomial test are performed for each model. These tests find that this condition is satisfied for all model presented in this investigation.

3.4. Model extensions

To obtain more information about the transmission mechanism from international monetary variables to the Japanese economy, we follow Kim (2001) in adding one variable at the time to the system presented in equations 1-4. These variables will include Japanese real exports, real imports and real trade balance and the real trade weighted index of the yen.¹²

To establish the robustness of results, the sensitivity of results to the lag selection strategy will be examined, and the implications of different contemporaneous restrictions for

¹¹ For more detail see Enders, *Applied Econometric*, p. 266.

¹² The real trade balance is defined by the Bank of Japan as real exports minus real imports denominated in 2010 prices, indexed so that for the base year 2010 the real trade balance is equal to 100.

the effect of international monetary variables on the other variables will be reported. Results are obtained and shown to be unchanged if a global commodity price index or a global energy price index are used instead of oil prices in the SVEC systems estimated. Extension of the model to allow for interaction of foreign and domestic variables and to recognize the global financial crisis will be estimated. A factor augmented VAR (FAVAR) to identify monetary policy shocks will also be estimated.

4. The empirical results

4.1. The impulse response effects of the structural monetary shocks

Figure 7 shows the dynamic response (or impulse response function) of the Japanese short term interest rate, M2, consumer price index, industrial production index, oil price index, and the nominal effect trade-weighted yen exchange rate, in the SVEC in equations (1)-(4) to impulse response to monetary shocks in China, the U.S. and the Euro area. The dashed lines represent a one standard error confidence band around the estimates of the coefficients of the impulse response functions.¹³ The first, second, and third columns show the responses of Japanese variables to structural innovations in China M2, Euro area M2, and U.S. M2, respectively.

In the first row in Figure 7, the short term Japanese interest rate rises significantly to innovations in China's M2. This response is consistent with the other effects of China's M2 that tend to be expansionary for the Japanese economy. The short term Japanese interest rate does not respond significantly to an innovation in Euro area M2. The short term Japanese interest rate falls significantly to innovations in the U.S. M2. This latter result is consistent with a defensive response by the Japan to a stimulus by the U.S. that weakens the dollar and makes Japanese goods more expensive in the U.S. In the second row of Figure 7, Japan's M2

¹³ The confidence bands are obtained using Monte Carlo integration as described by Sims (1980), where 5000 draws were used from the asymptotic distribution of the VAR coefficient.

tends to fall with positive shocks in M2 in China and in the U.S. and to rise with positive shocks in M2 Euro area. The effects of the international monetary shocks on Japan's M2 are not statistically significant.

In the third row in Figure 7, Japan's CPI is significantly affected by innovations in China's M2. A positive shock to China's M2 results in a rise in Japan's CPI that is statistically significant after three months. The magnitude of the rise in the CPI is gradually increasing. This effect on the CPI will be seen to be connected with the effect of shocks to China's M2 on oil and commodity prices. Japan's CPI is not significantly affected by shocks in Euro area M2. In Figure 7, a positive shock to U.S. M2 results in a fall in Japan's CPI that becomes statistically after seven months. These results on Japan's CPI will tie in with the effects of the international monetary shocks on global oil and commodity prices.

In the fourth row in Figure 7, Japan's industrial production is positively and significantly impacted by positive shocks to China's M2, not significantly impacted by shocks to Euro area M2, and negatively and significantly impacted by positive innovations in U.S. M2. These results will be related to the effect on Japan's real exports of shocks to each of the countries' monetary aggregates.

The effects of shocks to China's M2, U.S. M2, and Euro area M2 on international oil price are shown in the fifth row in Figure 7. A positive innovation in China's M2 has a statistically significant positive effect on oil prices that rapidly builds up over four or five months and then gradually increases in magnitude over 20 months.¹⁴ A positive shock in Euro area M2 has a positive effect on oil prices, but the effect is small and only marginally

¹⁴ The positive effect of China on oil prices in recent years has been established in the literature. Hamilton (2011) notes that China has a 6.3% compound annual growth rate for petroleum consumption since 1998 and argues that oil demand from emerging market countries has been the dominant influence on oil prices in recent years. Kilian and Hicks (2012) find that unexpected growth in emerging economies (primarily China and India) drove the increase in real oil price over 2003-2008. Ratti and Vespignani (2013) find that increased liquidity in China significantly raises real oil prices over 1996:1-2011:12, whereas increased liquidity in the G3 countries does not.

statistically significant. A shock in U.S. M2 has a statistically insignificant effect on global oil prices.

The response effects of the nominal trade-weighted Japanese yen exchange rate currency to shocks to China's M2, U.S. M2, and Euro area M2 are shown in the last row in Figure 7. The response to a positive shock to U.S. M2 is initially positive (the value of the yen is increased) but is not statistically significant. A positive shock in Euro area M2 has a negative effect on the nominal trade-weighted yen, but the effect is only statistically significant in the fourth month. A positive shock in China's M2 has a statistically significant negative effect on the nominal trade-weighted yen from impact. This effect would raise the price of imports in Japan and reinforce the effect on the Japanese consumer prices of the effect positive innovations in China's M2 on the U.S. dollar price of oil.

The significant effect of shocks to China's M2 on the nominal trade-weighted yen exchange rate requires further discussion. A rise in China's M2 facilitates domestic growth in China and increases China's demand for imports. The currencies of the countries supplying imports to China experience upward pressure. To stabilize the pegged exchange rate, China intervenes in the foreign exchange market and sells foreign currency. The net effect of these actions on the nominal trade-weighted yen exchange rate depends on the mix of imports, and the mix of foreign currencies sold in the foreign exchange market. The latter will be influenced by the weights assigned to currencies in the reference basket of major currencies to which the renminbi is allowed to float within a narrow margin. It is found as an empirical result that the net effect of China's actions following a positive innovation in China's M2 is a decrease in the nominal trade-weighted yen exchange rate.

4.2. Responses of Japan's trade variables to monetary shocks

The impact of monetary shocks originating in the U.S., China and the Euro area on Japanese trade variables will be now examined. Following a procedure in Kim (2001) the

Japanese real exports, real imports and real trade balance are now added one variable at a time as an additional variable in the SVEC system in equations (1)-(4). The first, second, and third columns in Figure 8 show the responses of Japanese variables to a impulse response in China M2, Euro area M2, and U.S. M2, respectively.

In the first and second rows of Figure 8 the impulse response functions of Japan's real exports and real imports to the international monetary shocks are reported. A positive shock to China's M2 significantly expands Japan's real exports. The rise in Japan's real exports builds up rapidly over five or six months and then persists. This result is consistent with increases in China's M2 being associated with significant reduction in the trade-weighted Japanese yen exchange rate. A positive shock to U.S. M2 has a significant negative effect on Japan's real exports. A positive shock in Euro area M2 does significantly impact Japan's real exports. In the second row of Figure 8, the effects of the international monetary shocks on Japan's real imports are small compared to the effects on Japan's real exports. The effects of shocks to China M2 and Euro area M2 on Japan's real imports are not statistically insignificant. A positive shock to U.S. M2 has a significant negative effect on Japan's real imports.

In the third row of Figure 8 the impulse response functions of Japan's real trade balance to structural innovations in China's M2, Euro area M2, and U.S. M2 are reported. Japan's real trade balance shows a statistically significant large positive response to positive shocks in China's M2. This is consistent with large positive increase in Japanese real exports attendant on depreciation of trade-weighted yen following a positive shock in China's M2 and with Japan's real imports being price inelastic. A positive shock in Euro area M2 also has a significant positive (but smaller) impact Japan's real trade balance. A positive shock to U.S. M2 has a significant negative effect on Japan's real trade balance.

In the last row of Figure 8, the responses of the real trade-weighted Japanese yen exchange rate currency to the international monetary shocks are reported. The results for the real trade-weighted Japanese yen to shocks are very similar to those for the nominal trade-weighted Japanese yen. A positive shock in China's M2 has a statistically significant negative effect on the real trade-weighted yen that persists. The response of the real trade-weighted Japanese yen to positive shocks to U.S. M2 Euro area M2 are not statistically significant.

5. Robustness checks

In this section several changes are made to the model in equations (1) through (5) to examine the robustness of results.

5.1. Lag length

First, we check the sensitivity of our results to the lag selection strategy. The Akaike Information Criterion (AIC) is also widely used in time series analysis when a longer structure is preferred; in our model this criterion selected three lags. Kim and Roubini (2000), Kim (2000) and Kim (2001) used a more arbitrary lags structure fixing the lags at six. Our model is robust to both the used of three lags selected by the AIC and six lags used by other researches.¹⁵

5.2. Contemporaneous restrictions

Secondly, we will use two different contemporaneous restrictions for the international monetary variables. The alternative SVEC systems are given by the following (based on equations (1) to (4)):

¹⁵ Although, we note that the error bands in the impulse response function increase when six lags are used.

$$X_t = \begin{bmatrix} 1 & -a_{01} & -a_{02} & -a_{03} & 0 & 0 & 0 & 0 & 0 \\ -a_{10} & 1 & -a_{12} & -a_{13} & 0 & 0 & 0 & 0 & 0 \\ -a_{20} & -a_{21} & 1 & -a_{22} & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 1 & -a_{34} & 0 & 0 & -a_{37} & -a_{38} \\ -a_{40} & -a_{41} & -a_{42} & -a_{43} & 1 & -a_{45} & -a_{46} & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & 1 & -a_{56} & -a_{57} & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 & 1 & -a_{67} & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & -a_{83} & -a_{84} & -a_{85} & -a_{86} & -a_{87} & 1 \end{bmatrix} \begin{pmatrix} \Delta \log(US M2_t) \\ \Delta \log(EU M2_t) \\ \Delta \log(China M2_t) \\ Japan IR_t \\ \Delta \log(Japan M2_t) \\ \Delta \log(Japan CPI_t) \\ \Delta \log(Japan IP_t) \\ \Delta \log(OP_t) \\ \Delta \log(TWI_t) \end{pmatrix} \quad (6)$$

$$X_t = \begin{bmatrix} 1 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 1 & 0 & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 1 & -a_{34} & 0 & 0 & -a_{37} & -a_{38} \\ 0 & 0 & 0 & -a_{43} & 1 & -a_{45} & -a_{46} & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & 1 & -a_{56} & -a_{57} & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 & 1 & -a_{67} & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 & 0 & 1 & 0 \\ -a_{80} & -a_{81} & -a_{82} & -a_{83} & -a_{84} & -a_{85} & -a_{86} & -a_{87} & 1 \end{bmatrix} \begin{pmatrix} \Delta \log(US M2_t) \\ \Delta \log(EU M2_t) \\ \Delta \log(China M2_t) \\ Japan IR_t \\ \Delta \log(Japan M2_t) \\ \Delta \log(Japan CPI_t) \\ \Delta \log(Japan IP_t) \\ \Delta \log(OP_t) \\ \Delta \log(TWI_t) \end{pmatrix} \quad (7)$$

The first alternative restriction scheme is presented in equation (6). In this equation we allow all monetary aggregates variables (international and domestic) to affect each other contemporaneously. In equation (7), we allow the international monetary variable to affect contemporaneously the Japanese TWI. Although, the log likelihood ratio test for over identified restrictions support the restrictions selected in equation (5) (with higher chi-square coefficient), our general results show to be robust to this two different restrictions sets. In Table 3, the log likelihood ratio for over-identification test Chi-square values are reported for the three models shown in equations (5), (6) and (7). The highest value for over-identification test restriction is for our model of choice in equation (5), indicating that the restriction cannot be rejected at a higher significance level than for the other models.

The estimation of the SVEC models outlined in equations (6) and (7) results in impulse response of the U.S. economic variables to positive shocks to China's M2 that are very similar to those shown in Figure 7. In all models a positive innovation in China's M2 is

associated with a positive and statistically significant effect on oil price (or on commodity prices) that builds up rapidly over three months and then persists for twenty month time horizon shown. Based on the models in equations (6) and (7), a positive innovation in China's M2 results in a positive and statistically significant effect on Japan's consumer price index that persists, a statistically significant depreciation of the nominal (and real) trade-weighted yen for about twelve months, and a positive and statistically significant effect on Japan's real trade balance for twenty month time horizon shown. These results are available from the authors.

5.3. Global commodity and energy price indices

In line with Dedola and Lippi (2005) and Anzuni *et. al.* (2012) measures of commodity price other than oil price are now introduced into the model. Results are essentially unchanged if a global commodity price index (or a global energy price index) are used instead of oil prices in the SVEC systems described in equations (1) to (5). In Figure 9 the response of global commodity price index (World Bank) and of the response of global energy price index (IMF) to the international monetary shocks are reported when these indices replace oil price in the equations (1) to (5). The results are similar to those for the response of oil price to the international monetary shocks.

A positive innovation in China's M2 has a statistically significant positive effect on global commodity prices that rapidly builds up over four or five months and then persists, and has a statistically significant positive effect on global energy prices that rapidly builds up over three months, persists and gradually increases in size. A positive shock in Euro area M2 has positive and statistically significant effects on global commodity and on global energy prices that persist. The effect of positive shock in Euro area M2 on global commodity prices and on global energy prices are smaller than the effect of positive innovations in China's M2 on global commodity and global energy prices. A shock in U.S. M2 has a statistically

significant effect on global commodity oil prices in the first month and the effect thereafter is statistically insignificant. The responses of global energy prices to shocks to U.S. M2 are not statistically significant.

5.4. Global financial crisis

We use monthly data from 1999 to 2012. Monetary policy has undoubtedly changed since the global financial crisis and there may be a structural change in the way monetary policy affects the economy. Overall, several aspects of macroeconomic variables had unique characteristics during the recent global recession. Perri and Quadrini (2011), for example, document an unprecedented business cycle synchronization among the developed countries during the last two quarters of 2008.¹⁶ China is not included in the analysis and did not experience the recession so marked among the developed countries. It would be useful to see if the results change considering a sample where the financial crisis is not included. To this end we will include the period of the global financial crisis

To deal with the global financial crisis we take several approaches. First, we introduce in a dummy variable that takes the value 1 from July 2008 to December 2008 and 0 otherwise into equation (1).¹⁷ Second, we estimate the model with truncated samples that exclude the worst of the crisis, that is over 1999:01 to 2008:05 and 2009:01 to 2012:12 and 1999:01 to 2008:08 and 2008:12 to 2012:12. Results are essentially unchanged from those in Figures 7 and 8 from following these strategies for dealing with the global financial crisis (and are

¹⁶ Perri and Quadrini (2011) find that the last two quarters of 2008 stands out in the post-war era in terms of GDP falling by so much and in all countries G7 countries. The G7 is the U.S., Japan, Germany, U.K., France, Italy and Canada. The correlation of GDP movements among the developed countries during the last two quarters of 2008 jumped from 0.3 to 0.7 and the sample standard deviation of the correlations fell from 0.19 to 0.09.

¹⁷ We follow Perri and Quadrini (2011) in identifying the third and fourth quarter months of 2008, July 2008 to December 2008, as being the global financial crises. In the monthly data other dates are possible. We experiment with September 2008 to November 2008 as being the global financial crises without changing results. Lehman Brothers filed for bankruptcy protection 15 September, and the global financial crises appears to have stabilized by the end of November 2008 with dramatic action by the US Federal Reserve pledging to purchase mortgage bonds guaranteed by Fannie Mae and Freddie Mac.

available from the authors). Results will be presented in the next subsection when taking account of the global financial crisis is combined with changes in model identification.

5.5. Interaction among foreign and domestic economies

In equation 1, foreign industrial productions and interest rates are taking as exogenous and consequently the model restricts the interaction among foreign industrial productions, foreign interest rates and those with the domestic economy. In this section we show that when both foreign industrial productions and interest rates are including in the model as endogenous variables, so as to allow for interaction of foreign and domestic variables, results are similar to those obtained in figure 7. To specify this alternative model we estimate a 15-endogenous-variable SVEC model where the U.S., Euro area and China interest rate are placed ahead of those countries M2 in equation 5 and these countries individual industrial productions are placed after foreign M2 in the same equation.¹⁸ In addition a dummy variable that takes the value 1 from July 2008 to December 2008 to global financial crisis.

Figure 10 shows the dynamic response (or impulse response function) of the Japanese variables in this SVEC model to impulse response to monetary shocks in the U.S., Euro area and China. Results with these additional endogenous variables and a global financial crisis variable are remarkably similar to those in Figure 7. The results for China's monetary expansion are unchanged. China's monetary expansion increases global oil prices and increases Japan's inflation, industrial production, exports (in results not shown) and decreases the nominal trade-weighted value of the yen. The results for the Euro area monetary expansion are modified somewhat, in that now the effects of innovations to Euro area M2 have statistically insignificant effects on the Japanese variables. The results for U.S. monetary expansion are little different from estimating the model (1) through (5).

5.6. The FAVAR model

¹⁸ Note that in this model, foreign variables only interact in the lag structure as the test for over identify restrictions select these variables as contemporaneously exogenous.

The identification strategy followed in this paper is inspired by Kim and Roubini (2000) and Kim (2001), among others. There has been a growing literature focused on the identification of monetary policy in a VAR framework.¹⁹ For example, Uhlig (2005) uses a methodology based on sign restrictions. In a related study, Bernanke et al. (2005) propose a Factor Augmented VAR (FAVAR) to identify monetary policy shocks. They find that a small number of variables in a VAR may not span the information set used by market participants, who are known to follow hundreds of data series. Therefore, a set of factors can summarize large amounts of information about an economy. Given that we are analyzing the effects of foreign shocks on the Japanese economy, it can be argued that there is no need to restrict the analysis to a small number of series given that we could use a large dataset with a methodology like the FAVAR model. Consequently, we estimate a FAVAR model to check the robustness of our results. To construct a FAVAR model we develop a principal component vector by including the variables: interest rate, consumer price index, industrial production, nominal trade weighted index for each of the other three largest economies (China, the U.S. and the Euro area). This vector is included as an endogenous variable and placed as the first variable in augmented system (1), (2), (3) and (5).²⁰ Results (available under request) are remarkably similar to those obtained in figure 10, showing the robustness of our model.

6. Discussion of the results

¹⁹ Another approach considered in the literature is use of a global VAR (GVAR) by Dees et al. (2007). The GVAR combines separate models for each of the many economies linking core variables within each economy with foreign variables using quarterly data. The foreign variables external to a domestic economy are trade-weighted. In our study we do not pursue a GVAR or GVEC for the following reasons. With monthly data, construction of time varying weights is not possible given the absence of bilateral trade data on a monthly basis for China and the fact that the activity variable in each country, industrial production, is an index. Also, the alternative of using fixed weights is also not appropriate given the rapid increase over time in the influence of China in the trade statistics.

²⁰ We allow this variable to interact contemporaneously with other variables in the system as indicated by the over identify log likelihood ratio test. The FAVAR approach in Bernanke et al. (2005) uses a two-step procedure suggested by Stock and Watson (2002). Factors are estimated by principal components before estimation of the factor-augmented VAR. The number of unobservable factors used for policy analysis may be selected on statistical grounds. We choose one factor and two lags to retain parsimony in the factor augmented VEC approach.

The predictions of the Mundell–Flemming and the sticky price (or sticky wage) intertemporal models provide a basis for the evaluation of expansionary monetary policy in the Euro area and the U.S. on Japan since their capital markets are fairly well integrated. In the Mundell–Flemming model monetary policy expansion results in beggar-thy-neighbour effects under flexible exchange rates. Monetary expansion in a country leads to currency depreciation and an improving trade balance through expenditure-switching, with the implication of falling incomes in other countries. The improvement in the trade balance and the fall in foreign income might be mitigated by the income-absorption effect (increased income in the inflating country leading to increased demand for imports).

In the intertemporal general equilibrium model of Obstfeld and Rogoff (1995), monetary expansion in a country leads to currency depreciation and also lowers the world real interest rate in proportion to the increase in the world money supply and this raises global demand.²¹ Currency depreciation of the country increasing the money supply shifts world demand toward goods production in that country at the expense of production in other countries. It is possible that output in the countries with appreciating currencies could fall, but if the inflating country is large and integrated into the global capital market the fall in world interest rates could have large stimulative effects on smaller economies that compensates for the expenditure switching effect.

Our empirical results suggest that U.S. monetary expansion leads in the first three or four months to an appreciation in the nominal trade-weighted yen (that is not statistically significant). There is a worsening in the Japanese real trade balance (the expenditure switching effect), with real exports falling more than real imports, and a fall industrial production. An income absorption effect (if any) does not offset expenditure switching. A statistically significant fall in the Japanese interest rate is associated with a positive shock in

²¹ Obstfeld and Rogoff (1995) build a general equilibrium two-country model of monopolistic competition and sticky prices in which the effects of monetary policy can be examined in welfare terms.

U.S. M2, but the impulse results overall suggest that U.S. monetary expansion has a depressing effect on the Japanese economy over the 1999-2012 period.

The results of U.S. monetary expansion for Japan can be contrasted with those obtained by Kim (2001) for the flexible exchange rate period 1974-1996.²² Kim (2001) finds that monetary expansion in the U.S. causes a reduction in interest rates and boom in the non-U.S. G-6 (in that GDP and industrial production increase in those countries). The trade balances of the non-U.S. G-6 with the rest of the world are not significantly affected by U.S. monetary expansion, indicating that income-absorption is important. Kim (2001) concludes that the boom in non-U.S. G-6 is due to the fact that the stimulative effect of the decline in non-U.S. G-6 interest rates attendant on U.S. monetary expansion.

The results for the effects of U.S. monetary expansion on Japan over 1999-2012 are very different from those reported for the period 1974-1996 during which positive U.S. monetary shocks have expansionary effects on the Japanese economy. There are several potential reasons for this difference. The U.S. economy is not as dominant as it once was and the Japanese economy has been in depressed state for an extended period. The potential for big effects on income through a fall in Japan's interest rate are precluded because Japan's interest rate is already at low levels.²³

Our empirical results suggest that Euro area monetary expansion does not have statistically significant effects on Japanese variables. Positive shocks to Euro area M2 are not significantly associated with movement in Japan's short-term interest rate, nominal trade-weighted yen, industrial production, inflation, real exports or real imports. Effects on trade are small from Euro area monetary expansion. In addition, prospective effects on Japan's

²² Kim (2001) finds that U.S. monetary policy expansion lead to a fall in the U.S. dollar. Income absorption effect of the monetary expansion is more important than expenditure substitution effect within the U.S. in that the U.S. trade balance worsens. Kim (2001) uses a VAR model with quarterly data.

²³ Ono (2006) argues that Japan's stagnation and decline in the marginal utility of consumption relative to that of liquidity is associated with appreciation of the yen relative to dollar, and decline in home output and increase in foreign output.

economy through Euro area monetary expansion influencing Japan's short-term interest rate are excluded because of the close to zero interest rates in Japan.

The discussion of the transmission mechanism of monetary shocks from the U.S. and the Euro area to Japan in terms of the Mundell–Flemming and the sticky price (or sticky wage) intertemporal models is about economies with flexible exchange rates that are integrated into world capital markets. These conditions do not necessarily apply to changes in China's money supply. China has a fixed exchange rate regime and maintains the value of the Chinese renminbi relative to the U.S. dollar and, since 2005, a basket of other currencies. China has extensive capital controls and the financial system remains tightly controlled. An interest rate mechanism is not available through capital markets even if Japan's interest rates were not at rock bottom levels. How are shocks to the Chinese money supply transmitted to Japan and the rest of the world?

The influence of China's monetary expansion on Japan is through the effect of China's growth on global oil and commodity prices, China's increased demand for imports, and China's exchange rate policy. China's monetary expansion is associated with statistically significant decreases in the nominal and real trade-weighted value of the yen. The significant rise in commodity prices in response to monetary expansion in China and decrease in the trade-weighted yen is reflected in significantly higher inflation in Japan. China's monetary expansion is associated with significant increases in Japan's real exports and industrial production.

7. Conclusions

It is found that over 1999:1-2012:12 China's monetary expansion has significant effects on Japan's economy that are quite different from those of the U.S. and Euro area. Capital controls in China provide room for monetary policy independence in China even

though the exchange rate is pegged (either to the U.S. dollar or to a basket of major currencies). Monetary expansion in China is reflected in significantly higher oil/commodity prices and inflation in Japan. In all models a positive innovation in China's M2 is associated with positive and statistically significant effects on oil price (or on commodity prices) and on Japan's consumer price index that build up rapidly over three or four months that then persist. China's monetary expansion significantly decreases the nominal and real trade-weighted yen exchange rates. China's monetary expansion is associated with significant increases in Japan's exports and industrial production.

In contrast, U.S. monetary expansion significantly lowers Japan's short-term interest rate. This result is in line with economic theory that monetary expansion within a large economy decreases real interest rates around the world. However, since Japan's interest rate is at such low levels over the period of analysis, large decreases in interest rates are not possible, precluding significant positive effects on income through an interest rate channel from monetary shocks in the U.S. (or the Euro zone). U.S. monetary expansion lowers Japan's exports and industrial production. Monetary expansion in the Euro area does not significantly affect Japan's exports, industrial production and interest rate. Japan's real trade balance improves with monetary expansion in China and worsens with monetary expansion in the U.S. A structural vector error correction model is used to examine international monetary transmission to Japan from U.S., China and the Euro area. Results are robust across differently specified structural vector error correction models.

References

- Anzuini, A., Lombardi, M.J., Pagano, P., 2012. The impact of monetary policy shocks on commodity prices. Bank of Italy Working Paper, Number 851.
- Bachmeier, L.J., Swanson, N.R., 2005. Predicting Inflation: Does The Quantity Theory Help? *Economic Inquiry* 43, 570-585.
- Bernanke, B., Boivin, J., Elias, P.S., 2005. Measuring the Effects of Monetary Policy: A Factor-augmented Vector Autoregressive (FAVAR) Approach. *Quarterly Journal of Economics* 120, 387-422.
- Cai, Z., Chen, L. and Fang, Y., 2012, A new forecasting model for USD/CNY exchange rate", *Studies in Nonlinear Dynamics & Econometrics*, 16(2), forthcoming.
- Canova, F., 2005. The transmission of US shocks to Latin America. *Journal of Applied Econometrics* 20, 229–251.
- Christiano, L.J., Eichenbaum, M., Evans, C., 1999. Monetary policy shocks: What have we learned and to what end? In: Taylor, J.B., Woodford, M. (Eds.), *Handbook of Macroeconomics*, Vol. 1A. North-Holland, Amsterdam, pp. 65–148.
- Dedola, L., Lippi, F., 2005. The monetary transmission mechanism: Evidence from the industries of five OECD countries, *European Economic Review* 49, 1543-1569.
- Dees, S., di Mauro, F., Pesaran, M.H., Smith, L.V., 2007, Exploring the international linkages of the euro area: a global VAR analysis. *Journal of Applied Econometrics* 22, 1-38.
- Di Giovanni, J., Shambaugh, J. C., 2008. The impact of foreign interest rates on the economy: The role of the exchange rate regime. *Journal of International Economics* 74, 341-361.
- Fan, L., Yu, Y., Zhang, C., 2011. An empirical evaluation of China's monetary policies. *Journal of Macroeconomics* 33, 358–371.
- Fang, Y., Huang, S., Niu, L., 2012. De facto currency baskets of China and East Asian economies: The rising weights. Bank of Finland BOFIT Discussion Papers, 2-2012.
- Fujiwara, I., 2007. Is there a direct effect of money?. Money's role in an estimated monetary business cycle model of the Japanese economy. *Japan and the World Economy* 19, 329-337.
- Garratt, A., Koop, G., Mise, E., Vahey, S.P., 2009. Real-Time Prediction With U.K. Monetary Aggregates in the Presence of Model Uncertainty. *Journal of Business & Economic Statistics*, American Statistical Association 27, 480-491.
- Goodfriend, M., Prasad, E., 2007. A framework for independent monetary policy in China. *CESifo Economic Studies* 53, 2-41.
- Gordon, D.B., Leeper, E.M., 1994. The Dynamic Impacts of Monetary Policy: An Exercise in Tentative Identification. *Journal of Political Economy* 102, 1228-47.

- Grilli, V., Roubini, N., 1995. Liquidity and exchange rates: puzzling evidence from the G-7 countries. Working paper, Yale University, CT.
- Hamilton, J.D., 2011. Historical Oil Shocks, NBER Working Paper No. 16790. Forthcoming in *The Routledge Handbook of Major Events in Economic History*, Randall E. Parker (Author, Editor), Robert M. Whaples (Editor), Routledge, 2013.
- Hosono, K., 2006. The transmission mechanism of monetary policy in Japan: Evidence from banks' balance sheets. *Journal of the Japanese and International Economies* 20, 380-405.
- Johansson, A.C., 2012. China's Growing Influence in Southeast Asia – Monetary Policy and Equity Markets. *The World Economy* 35, 816-837.
- Kilian, L., and B. Hicks (2012), "Did Unexpectedly Strong Economic Growth Cause the Oil Price Shock of 2003-2008?" *Journal of Forecasting*, forthcoming.
- Kim, S., 1999. Do monetary policy shocks matter in the G-7 countries? Using common identifying assumptions about monetary policy across countries. *Journal of International Economics* 48, 387-412.
- Kim, S., 2001. International transmission of US Monetary policy shocks: evidence from VARs. *Journal of Monetary Economics* 48, 339-372.
- ~~Kim, S., 2001b. Effects of monetary policy shocks on the trade balance in small open European countries. *Economics Letters* 71, 197-203.~~
- Kim, S., Roubini, N., 2000. Exchange rate anomalies in the industrial countries: a solution with a structural VAR approach. *Journal of Monetary Economics* 45, 561-586.
- ~~Kim, S., Yang, D.Y., 2012. International monetary transmission in East Asia: Floaters, non-floaters, and capital controls. *Japan and the World Economy* 24, 305-316.~~
- Koray, F., McMillin, W. D., 1999. Monetary shocks, the exchange rate, and the trade balance. *Journal of International Money and Finance* 18, 925-940.
- Kozłuk, T., Mehrotra, A., 2009. The Impact of Chinese Monetary Policy Shocks on East and South-East Asia. *Economics of Transition* 17, 121-45.
- Ma, G, McCauley, R.N., 2011. The evolving renminbi regime and implications for Asian currency stability. *Journal of the Japanese and International Economies* 25, 23-38.
- Nagayasu, J., 2007. Empirical analysis of the exchange rate channel in Japan. *Journal of International Money and Finance* 26, 887-904.
- Nakashima, K., Saito, M., 2012. On the comparison of alternative specifications for money demand: The case of extremely low interest rate regimes in Japan. *Journal of the Japanese and International Economies* 26, 454-471.

- Obstfeld, M., Rogoff, K., 1995. Exchange rate dynamics redux. *Journal of Political Economy* 103, 624–660.
- Ono, Y., 2006. International asymmetry in business activity and appreciation of a stagnant country's currency. *The Japanese Economic Review* 57, 101-120.
- Pesaran, M.H., Schuermann, T., Weiner, S.M. 2004. Modeling regional interdependencies using a global error-correcting macroeconomic model. *Journal of Business and Economic Statistics* 22, 129-162.
- Perri, F., Quadrini, V., 2011. International Recessions. National Bureau of Economic Research, NBER Working Paper 17201.
- Ratti, R.A., Vespignani, J., 2013. Liquidity and Crude Oil Prices: China's Influence Over 1997-2011. *Economic Modelling* 33, 517-525.
- Sims, C.A., 1980. Macroeconomics and Reality. *Econometrica* 48, 1-48.
- Sims, C.A., Zha, T., 1995. Does monetary policy generate recessions?: Using less aggregate price data to identify monetary policy. Working paper, Yale University, CT.
- Sims, C.A., Zha, T., 2006. Does Monetary Policy Generate Recessions?, *Macroeconomic Dynamics* 10, 231-272.
- Stock, J.H., Watson, M.W., 2002. Forecasting using principal components from a large number of predictors. *Journal of the American Statistical Association* 97, 1167–1179.
- Sun, H., 2009. Autonomy and Effectiveness of Chinese Monetary Policy under the De Facto Fixed Exchange Rate System. *China and World Economy* 17, 23-38.
- Svensson, L., Van Wijnbergen, S., 1989. Excess capacity, monopolistic competition and international transmission of monetary disturbances. *Economic Journal* 99, 785–805.
- Swanson, N.R., 1998. Money and output viewed through a rolling window. *Journal of Monetary Economics* 41, 455-474.
- Uhlig, H., 2005. What are the effects of monetary policy on output? Results from an agnostic identification procedure. *Journal of Monetary Economics* 52, 381-419.

Table 1: Test for unit roots 1999:1-2012:12:

Endogenous variables					
Variables	ADF	KPSS	First difference	ADF	KPSS
$\log(US M2_t)$	0.20	1.79***	$\Delta\log(U.S.M2_t)$	-10.08***	0.09
$\log(China M2_t)$	0.37	1.62***	$\Delta\log(China M2_t)$	-2.94**	0.73*
$\log(EU M2_t)$	0.30	1.24***	$\Delta\log(EU M2_t)$	-12.86***	0.25
$\log(JAP M2_t)$	0.33	1.35***	$\Delta\log(U.S.M2_t)$	-12.98***	0.24
$\log(JAP CPI_t)$	-2.24	0.94***	$\Delta\log(U.S.CPI_t)$	-2.76***	0.10
$\log(JAP IP_t)$	-2.52	0.35*	$\Delta\log(U.S.IP_t)$	-10.36***	0.06
$\log(OP_t)$	-2.43	1.48***	$\Delta\log(OP_t)$	-9.86***	0.10
$\log(JAP NTWI_t)$	-1.71	0.58**	$\Delta\log(JAP TWI_t)$	-9.65***	0.08

Notes: The variables are U.S M2, China M2, Euro area M2 (EU M2), Japan M2 (JAP M2), Japan consumer price index (JAP CPI), Japan industrial production (JAP IP), global oil price (OP), and nominal exchange rate Japanese foreign exchange rate (foreign/domestic) (JAP NTWI). The null hypothesis for the Augmented Dickey-Fuller (ADF) test is the variable has a unit root and the null hypothesis for the Kwiatkowski–Phillips–Schmidt–Shin (KPSS) test is the variable is stationary. The first difference of the series is indicated by Δ . The lag selection criteria for the ADF is based on Schwarz information Criteria (SIC) and for the KPSS is the Newey-West Bandwidth. ***, **, * indicates rejection of the null hypothesis at 1%, 5% and 10%, levels of significance.

Table 2: Cointegration test

Variables: $\log(Japan IP_t)$, $\log(Japan M2_t)$, $\log(Japan CPI_t)$					
Selected (0.05 level*) Number of Cointegrating Relations by Model					
Data Trend:	None	None	Linear	Linear	Quadratic
Test Type	No Intercept	Intercept	Intercept	Intercept	Intercept
	No Trend	No Trend	No Trend	Trend	Trend
Trace	1	1	1	1	1
Max-Eig	1	1	1	1	1

*Critical values based on MacKinnon-Haug-Michelis (1999)

Table 3: Log likelihood ratio test for over-identification restrictions

Null Hypothesis : Restrictions are valid

Model restrictions/Equations	5	6	7
Chi-square value	77.9	42.6	30.1

Notes: Log likelihood ratio for over-identification test Chi-square values are reported for each of the three models shown in equations (5), (6), (7). The test is for non-recursive identification restrictions in the contemporaneous matrix restrictions in equations, based on equations (1) through (4). The highest value for over-identification test restriction is for model of choice in equation (5), indicating that the restriction cannot be rejected at higher significant level than for the other models.

Figure 1: M2 in billions of USD for the U.S., EU, Japan and China

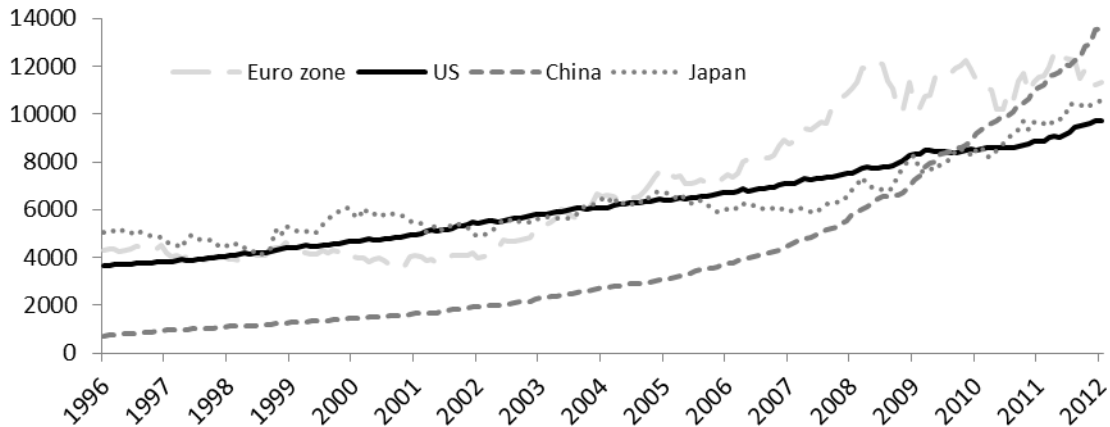
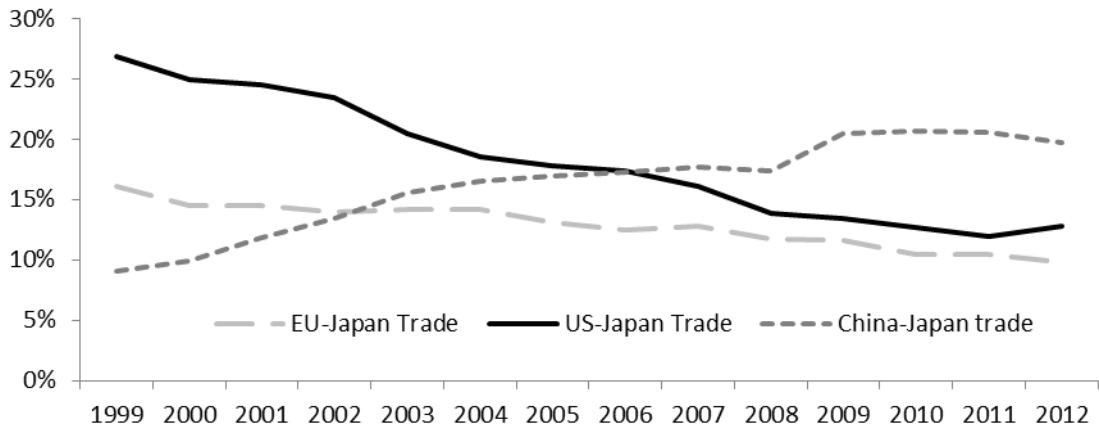
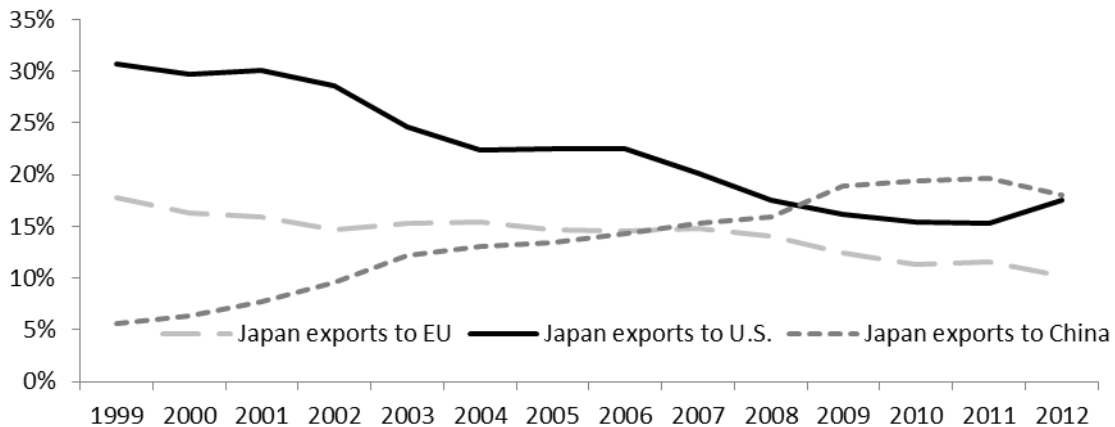


Figure 2: Economy region/country trade (exports plus imports) as percentage % of total trade to Japan



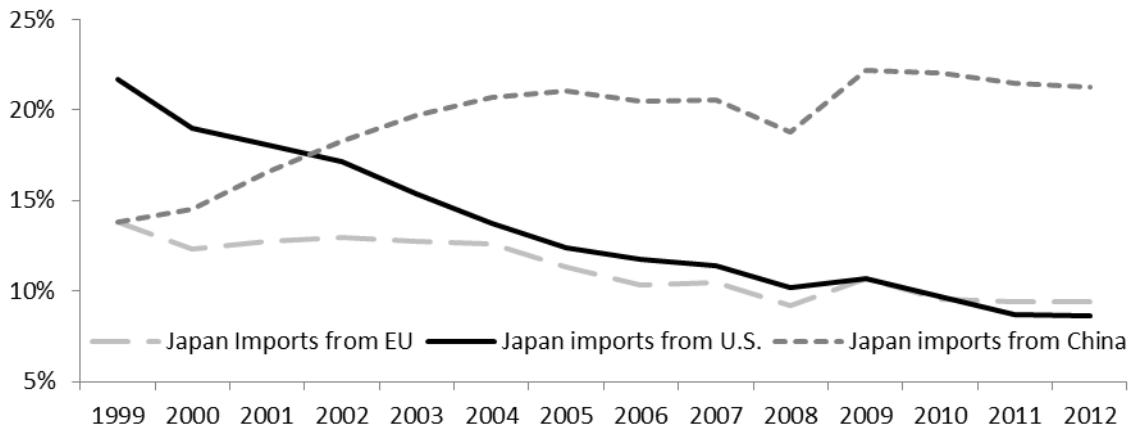
Source: Trade statistics of Japan (Minister of finance).

Figure 3: Economy region/country exports as percentage % of total trade to Japan



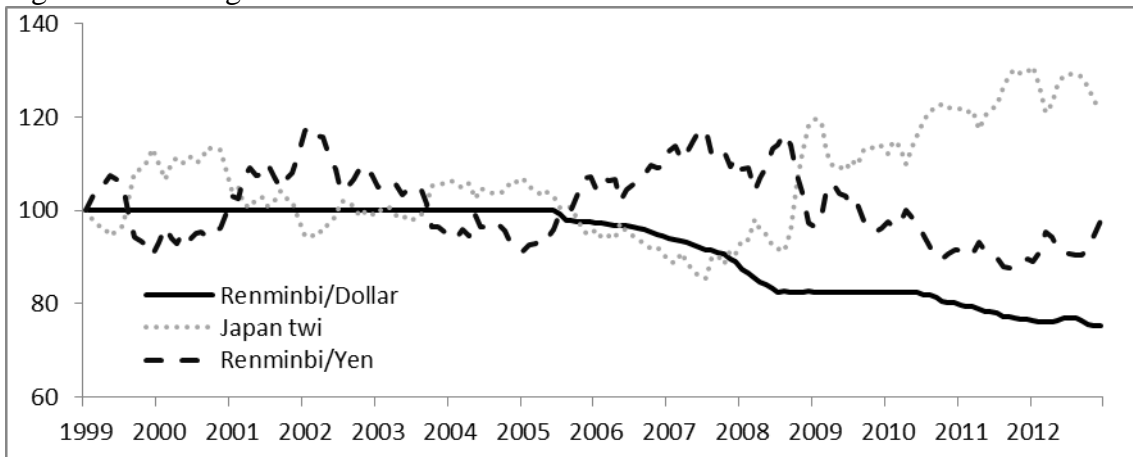
Source: Trade statistics of Japan (Minister of finance).

Figure 4: Economy region/country imports s as percentage % of total trade to Japan



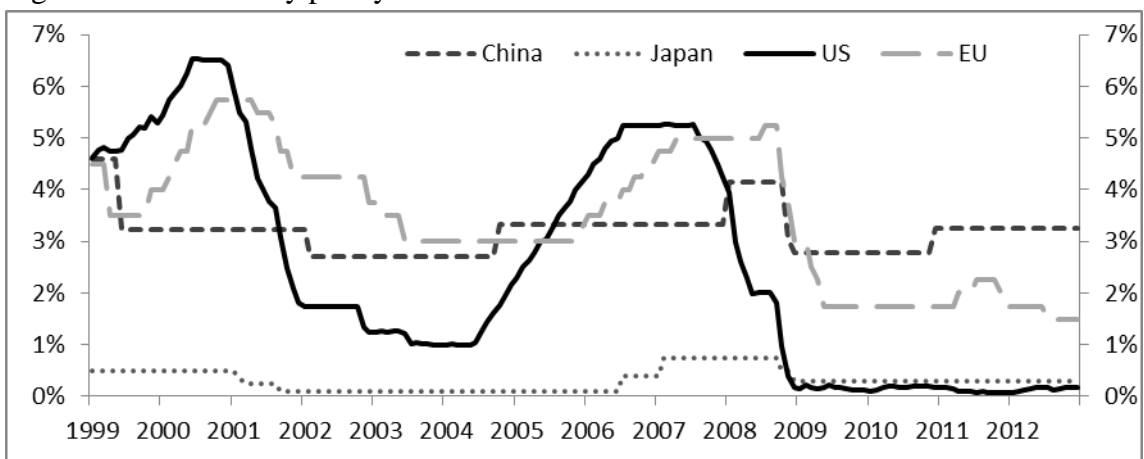
Source: Trade statistics of Japan (Minister of finance).

Figure 5: Exchange rates



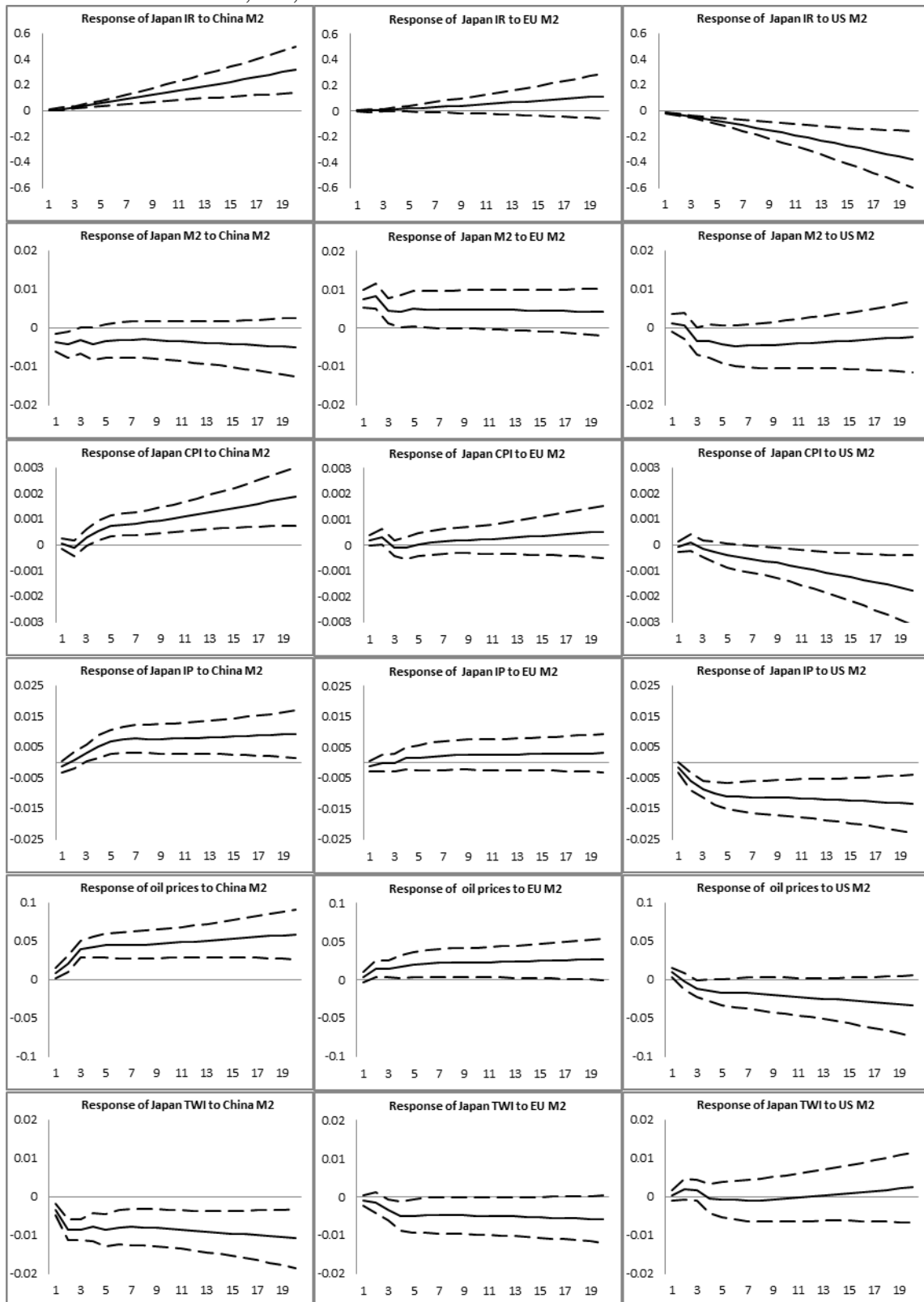
Source: Federal Reserve of St. Louis (FRED).

Figure 7: G4 Monetary policy interest rates



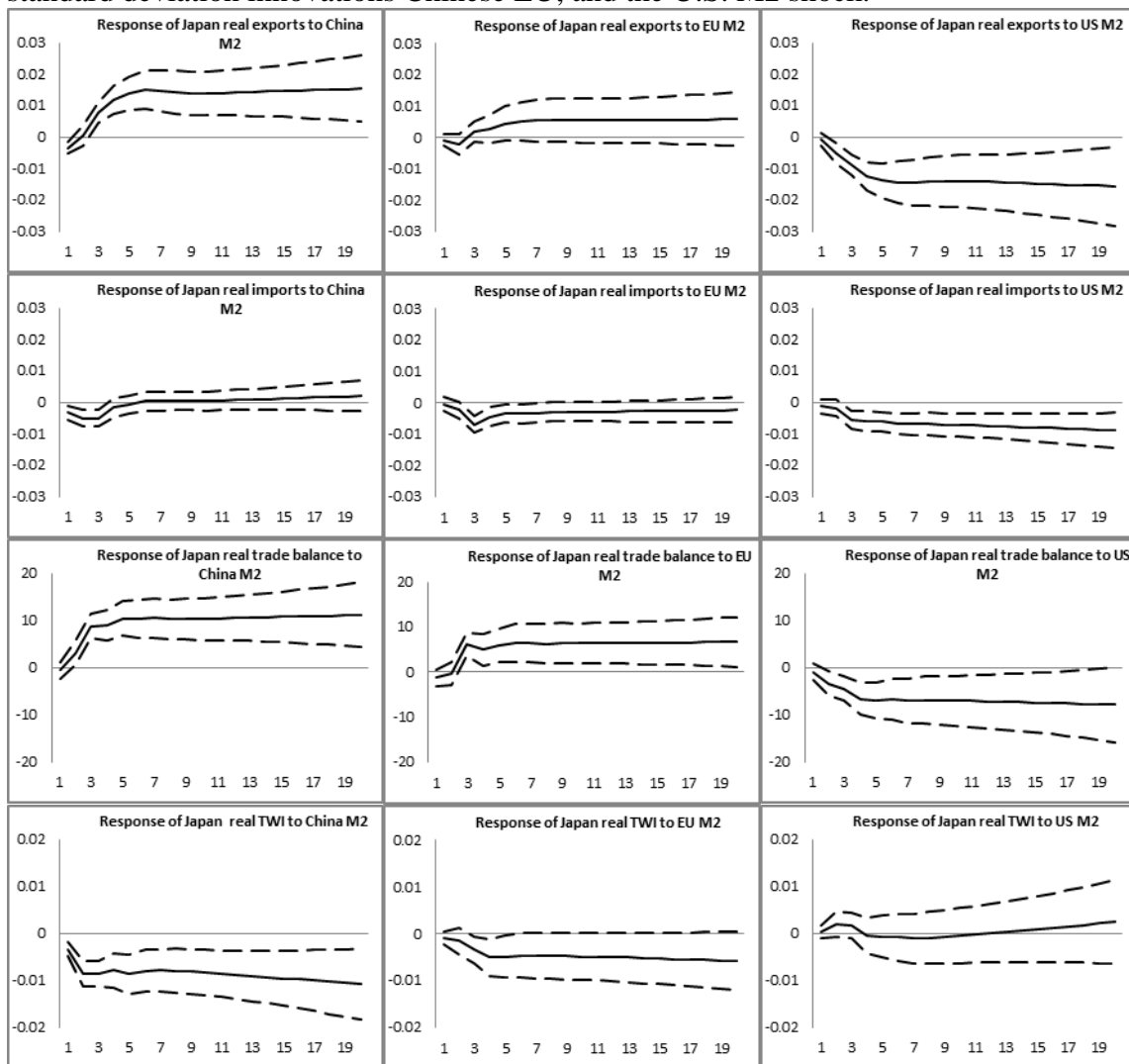
Source: Federal Reserve of St. Louis (FRED). The interest rates are the Federal Funds rate for the U.S. and the central bank discount rates for China, Japan, and the Euro area.

Figure 7: Impulse responses of Japanese economic variables to one standard deviation innovations of Chinese, EU, and the U.S. M2 shocks.



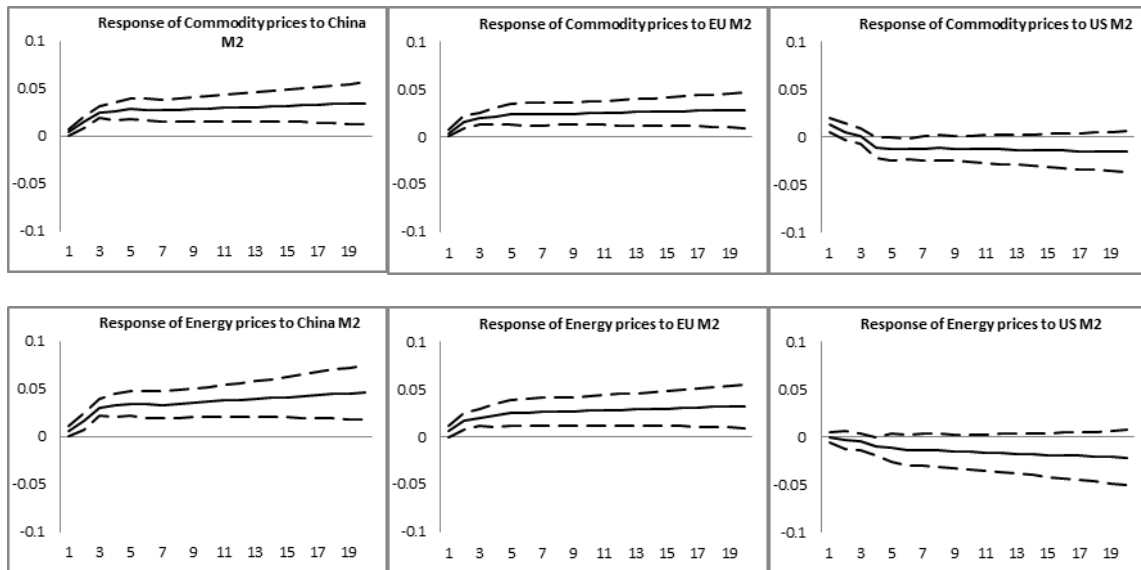
Notes: The confidence bands are obtained using Monte Carlo integration as described by Sims (1980), where 5000 draws were used from the asymptotic distribution of the VAR coefficient. The variables are U.S. M2, China M2, Euro area M2 (EU M2), Japan M2, Japanese industrial production (Japan IP), Japanese consumer price index (Japan CPI), oil prices, short term Japanese interest rate (Japan IR), nominal trade-weighted Japanese currency foreign exchange rate (Japan TWI).

Figure 8: Model extensions. Impulse responses of Japanese RTWI, NX, RX and RTB to one standard deviation innovations Chinese EU, and the U.S. M2 shock.



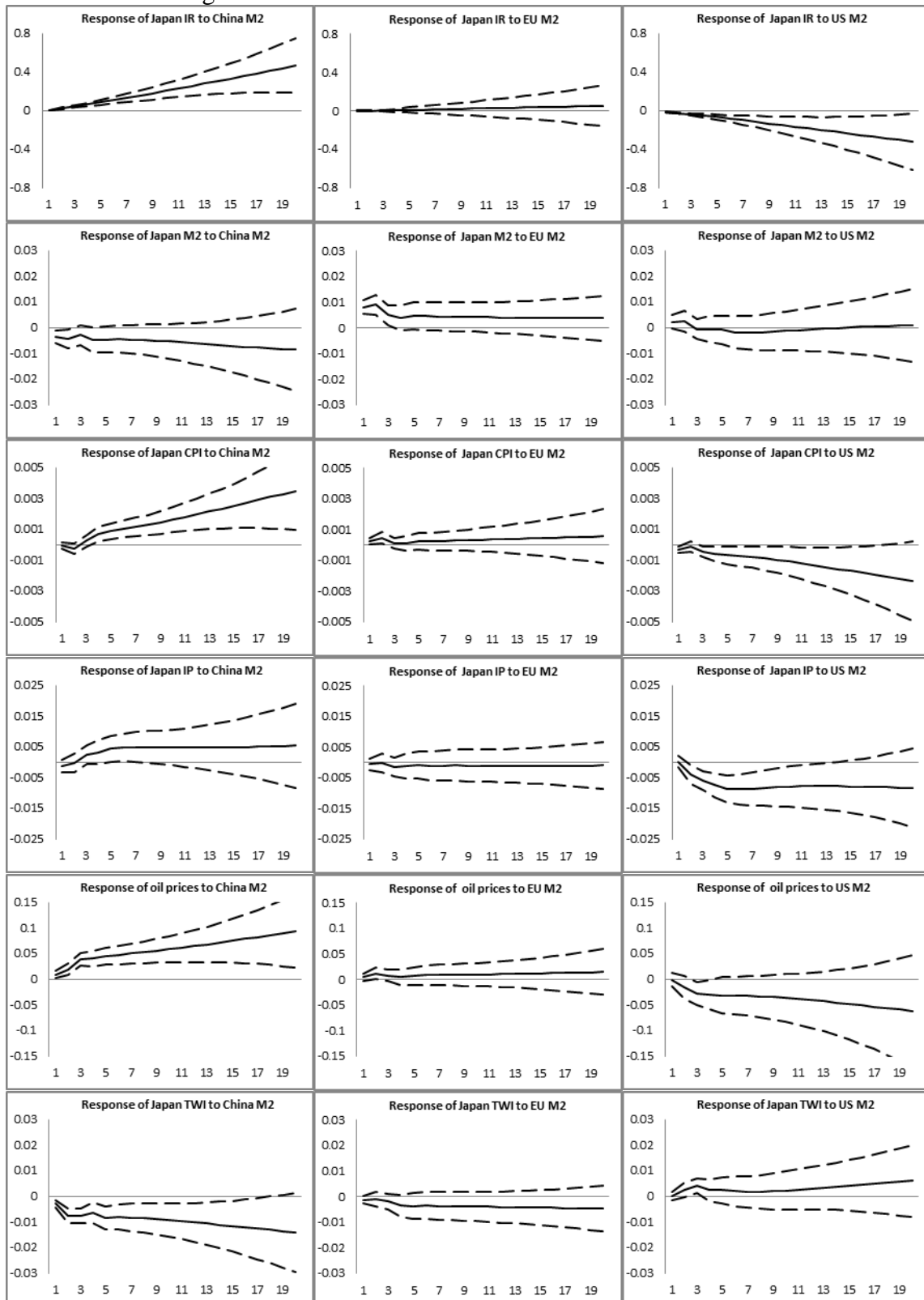
Notes: The confidence bands are obtained using Monte Carlo integration as described by Sims (1980), where 5000 draws were used from the asymptotic distribution of the VAR coefficient. The variables are U.S. M2, China M2, Euro area M2 (EU M2), Japan M2, Japanese industrial production (Japan IP), Japanese consumer price index (Japan CPI), oil prices, short term Japanese interest rate (Japan IR), real trade-weighted Japanese currency foreign exchange rate (Japan TWI), Japanese nominal (NX) and real exports (RX) and Japanese real trade balance (RTB).

Figure 9: Response of commodity and energy prices to monetary shocks: 1999:01-2012:12



Notes: In Figure 9 the response of all commodity price index (World Bank) and the response of energy price index (IMF) to the international monetary shocks are reported when these indices replace oil price in the SVEC system (1) to (5).

Figure 10: Robustness analysis allowing for interactions between domestic and foreign variables and with global financial crisis variable



Notes: Figure 10 shows the impulse response of the Japanese variables in the SVEC model, that allows for interaction between domestic and foreign variables and includes a global financial crisis variable taking the value 1 from July 2008 to December 2008 and 0 otherwise, to monetary shocks in China, the U.S. and the Euro area.